# THE MONTY HALL PROBLEM

THE MONTY HALL PROBLEM

$$
\begin{array}{|c|c|c|}\n\hline\nA & B & C \\
\hline\nD & D & D \\
\hline\nT & D & D\n\end{array}
$$



Roll 2 dice...





Roll 2 dice...  $P(A) = P(sum = 8)$  $P(B) = P(both are even)$  $\begin{picture}(180,10) \put(0,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}} \put(15,0){\line(1,0){155}}$ If we knew that both are even, then what is the probability that the sum is 8?



Roll 2 dice ... 
$$
P(A) = P(sum = 8)
$$
 "prob. A given B"  
\na. 3  
\n $P(B) = P(both are even)$   
\n $P(A | B)$   
\n $P(A | B)$ 

$$
\mathcal{L} = \mathcal{L} \left( \mathcal{L} \right)
$$

Roll 2 dice ... 
$$
P(A) = P(svm = 8)
$$
 "prob. A given B"  
\n $P(A | B)$   
\n $P(B) = P(both are even)$   
\n $l^{\{}$  we knew that both are even, then what is the probability that the sum is 8?

Roll 2 dice ... 
$$
P(A) = P(sum = 8)
$$
 "prob. A given B"  
\na. 1  
\n $P(A | B)$   
\n $P(B) = P(both are even)$   
\nIf we knew that both are even, then what is the probability that the sum is 8?  
\n $A = \{(2,6), (3,5), (4,4), (5,3), (6,2)\}$   
\n $B = \{(2,2), (2,4), (2,6), (4,2), (4,4), (4,6), (6,2), (6,4), (6,6)\}$ 

Roll 2 dice ... 
$$
P(A) = P(sum = 8)
$$
 "prob. A given B"  
\na. P(B) = P(both are even)  
\n $P(A | B)$   
\n $P(B) = P(both are even)$   
\n $P(A | B)$   
\n $P(A | B)$ 

Roll 2 dice ... 
$$
P(A) = P(sum = 8)
$$
  
\n $\frac{a}{\therefore 1}$   $P(B) = P(both \text{ are even})$   
\n $\frac{P(A \mid B)}{\therefore 1}$   
\n $\frac{P(B)}{\therefore 1}$   $P(B) = P(both \text{ are even})$   
\n $\frac{P(A \mid B)}{\therefore 1}$   
\n $P(B \mid B)$   
\n $P(A \mid B)$   
\n $P(A \mid B)$   
\n $P(A \mid B)$   
\n $P(A \mid B)$   
\n $P(B) = \frac{3}{9}$   
\n $B = \{(2,2), (2,4), (2,6), (4,2), (4,4), (4,6), (6,2), (6,4), (6,6)\} \neq P(A)$   
\n $\frac{P(A \mid B) = \frac{3}{9}}{\frac{7}{9}}$   
\n $\frac{P(A \mid B) = \frac{3}{9}}{\frac{7}{9}}$ 

$$
P(A) = P(sum = 8)
$$
  
 $P(B) = P(both are even)$ 



$$
P(A) = P(som = 8)
$$
  
 $P(B) = P(both are even)$ 



$$
P(A) = P(sum = 8)
$$
  
\n $P(B) = P(both are even)$   
\n $P(A) = \frac{9}{600} \text{ area } 5$   
\n $P(A) = \frac{9}{600} \text{ area } 36$ 





When we establish  $B$  then the universe shrinks.



When we establish  $B$  then the universe shrinks. The probability that A holds is normalized: <u>remaining valid green area</u>



When we establish  $B$  then the universe shrinks. The probability that A holds is normalized: <u>remaining valid green area</u>

$$
P(A|B) = ?
$$



When we establish  $B$  then the universe shrinks. The probability that A holds is normalized: <u>remaining valid green area</u>

$$
P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{3}{3+6}
$$

another example 
$$
Flip \alpha coin 5 times. P(1st flip = T) = \frac{1}{2}
$$

another example. Flip a coin 5 times. 
$$
P(1st flip = T) = \frac{1}{2}
$$
  
But what if you know that 3 of the 5 flips were H ?

another example: Filip a coin 5 times. 
$$
P(1s+1|_{p} = T) = \frac{1}{2}
$$
  
But what if you know that 3 of the 5 flips were H ?  
  
 $\downarrow$  P(1s+1|\_{p} = T | 3\cdot H) = ?

 $\bullet$ 

another example: Filip a coin 5 times. 
$$
P(1s+1|_{P} = T) = \frac{1}{2}
$$
  
But what if you know that 3 of the 5 flips were H ?  
  
 $P(1s+1|_{P} = T | 3\cdot H) = \frac{P[(1s+1|_{P} = T) \cap (3\cdot H)]}{P(3\cdot H)}$ 

Another example: 
$$
Flip \propto \text{coin 5 times. } P(1s + flip = T) = \frac{1}{a}
$$

\nBut what if you know that 3 of the 5 flips were H?

\n
$$
P(1s + flip = T | 3 \cdot H) = \frac{P[(1s + flip = T) \cap (3 \cdot H)]}{P(3 \cdot H)}
$$

$$
P(3\cdot H) = ?
$$

Another example: Fig a coin 5 times. 
$$
P(1s+1|_{P} = T) = \frac{1}{2}
$$

\nBut what if you know that 3 of the 5 flips were H ?

\n
$$
P(1s+1|_{P} = T | 3\cdot H) = \frac{P[(1s+1|_{P} = T) \cap (3\cdot H)]}{P(3\cdot H)}
$$
\n
$$
P(3\cdot H) = \frac{\binom{5}{3}}{2^{5}} \Rightarrow \text{Sample space.}
$$

Another example: 
$$
Flip \propto \text{coin 5 times. } P(1s + P1; p = T) = \frac{1}{2}
$$

\nBut what if you know that 3 of the 5 flips were H?

\n
$$
P(1s + P1; p = T | 3 \cdot H) = \frac{P[(1s + P1; p = T) \cap (3 \cdot H)]}{P(3 \cdot H)}
$$

\n
$$
P(3 \cdot H) = \frac{\left(\frac{5}{3}\right)}{2^5} \Rightarrow \frac{\text{Ways to chose}}{3 \text{ sample space.}} = \frac{\frac{5!}{3!2!}}{32} = \frac{5}{16}
$$

Another example: Fig a coin 5 times. 
$$
P(1s+1|p = T) = \frac{1}{a}
$$

\nBut what if you know that 3 of the 5 flips were H ?

\n
$$
P(1s+1|p = T | 3\cdot H) = \frac{P[(1s+1|p = T) \cap (3\cdot H)]}{P(3\cdot H)}
$$
\n
$$
P(3\cdot H) = \frac{\left(\frac{5}{3}\right)}{2^5} \Rightarrow \frac{Ways \text{ to choose}}{3 \text{ points for H.}} = \frac{\frac{5!}{3!2!}}{32} = \frac{5}{16}
$$

$$
P[(1_{s}f \ell | i_{p} = T) \cap (3 \cdot H)] = ?
$$

Another example: Fig a coin 5 times. 
$$
P(1sf flip = T) = \frac{1}{2}
$$

\nBut what if you know that 3 of the 5 flips were H ?

\n
$$
P(1sfflip = T | 3 \cdot H) = \frac{P[(1sfflip = T) \cdot (3 \cdot H)]}{P(3 \cdot H)}
$$
\n
$$
P(3 \cdot H) = \frac{\frac{5!}{3!}}{2^5} \Rightarrow \frac{Ways \cdot b \cdot c \cdot b \cdot c \cdot e}{3 \cdot 2^5} = \frac{\frac{5!}{3!2!}}{32} = \frac{5}{16}
$$

\n
$$
P[(1sfflip = T) \cdot (3 \cdot H)] = \frac{1}{T} \cdot HHH
$$
\n
$$
P[(1sfflip = T) \cdot (3 \cdot H)] = \frac{1}{T} \cdot HHH
$$

\n
$$
T \cdot HHH
$$

\n
$$
T \cdot HHH
$$

\n
$$
T \cdot HHH
$$

Another example: 
$$
Flip \, \alpha \, \text{coin 5 times. } P(1s + f1; p = T) = \frac{1}{2}
$$

\nBut what if  $\text{yon know that 3 of the 5 f1} \text{ is } \text{were } H$ ?

\n
$$
L \rightarrow P(1s + f1; p = T | 3 \cdot H) = \frac{P[(1s + f1; p = T) \cap (3 \cdot H)]}{P(3 \cdot H)}
$$
\n
$$
P(3 \cdot H) = \frac{\left(\frac{5}{3}\right)}{2^5} \rightarrow \frac{Ways}{3} \text{ points for } H = \frac{5!}{3!2!} = \frac{5}{16}
$$

\n
$$
P[(1s + f1; p = T) \cap (3 \cdot H)] : T \text{ HHTH} \qquad \frac{4}{32} \quad \text{OR } \frac{1}{2} \cdot \frac{\left(\frac{4}{3}\right)}{2^4} = \frac{1}{8}
$$
\n
$$
T \text{ HHTH} \qquad \frac{4}{32} \quad \text{OR } \frac{1}{2} \cdot \frac{\left(\frac{4}{3}\right)}{2^4} = \frac{1}{8}
$$

Another example: Filip a coin 5 times. 
$$
P(1sf flip = T) = \frac{1}{2}
$$

\nBut what if you know that 3 of the 5 flips were H ?

\nHint of having a biased coin : 60-40 vs 50-50 then the first flip has 40% for T  $\rightarrow \frac{2}{5}$ 

 $\overline{\mathbf{r}}$ 

=  $P(A)$   $\cap$   $P(B|A)$   $\cap$   $P(C|(A \cap B))$  ...

$$
Flip \alpha coin x3 : P(3rd=T|1st=H) = ?
$$
$$
F\left(\frac{1}{p} \alpha \text{ coin } x \frac{3}{3} : P\left(3x \frac{1}{5} + 1\right) =
$$
  
=  $\frac{P[(3x \frac{1}{5} + 1)]}{P(15 \frac{1}{5} + 1)}$ 

$$
Flip \alpha \text{ coin } x3 : P(3rd = T | 1st = H) =
$$
  
=  $\frac{P[(3rd = T) \cap (1st = H)]}{P(1st = H)} = \frac{\frac{2}{8} \rightarrow \text{sample space}}{1/2} = \frac{1}{2}$ 

$$
Flip \alpha \text{ coin } x3 : P(3rd = T | 1st = H) =
$$
  
=  $\frac{P[(3rd = T) \cap (1st = H)]}{P(1st = H)} = \frac{\frac{2}{8} \longrightarrow \text{sample space}}{V_2} = \frac{1}{2}$ 

$$
Notice P(3rd=T) = \frac{1}{2}
$$

$$
Flip \alpha \text{ coin } x3 : P(3rd = T | 1st = H) =
$$
\n
$$
= \frac{P[(3rd = T) \cap (1st = H)]}{P(1st = H)} = \frac{\frac{2}{8} \longrightarrow \text{sample space}}{V_2} = \frac{1}{2}
$$

Notice P(3rd=T) = = so knowledge of (1st=H) was useless.

INDEPENDENCE

$$
Flip \alpha \text{ coin } x3 : P(3rd = T | 1st = H) =
$$
\n
$$
= \frac{P[(3rd = T) \cap (1st = H)]}{P(1st = H)} = \frac{\frac{2}{8} \longrightarrow \text{sample space}}{\frac{1}{2}} = \frac{1}{2}
$$

Notice 
$$
P(3rd=T) = \frac{1}{2}
$$
 so knowledge of  $(1st=H)$  was useless.  
A & B are independent if  $P(A) = P(A|B)$ 

INDEPENDENCE

$$
Flip \alpha \text{ coin } x3 : P(3rd = T | 1st = H) =
$$
\n
$$
= \frac{P[(3rd = T) \cap (1st = H)]}{P(1st = H)} = \frac{\frac{2}{8} \longrightarrow \text{sample space}}{V_2} = \frac{1}{2}
$$

Notice 
$$
P(3rd=T) = \frac{1}{2}
$$
 so knowledge of  $(1st=H)$  was useless.  
A & B are independent if  $P(A) = P(A|B)$   
:  $P(B) = P(B|A)$  [equivalent]

$$
P(A|B) = \frac{P(A \cap B)}{P(B)}
$$











 $P(A) =$ 



$$
4: 1st person is a girl
$$
\n
$$
B: 2nd person is a girl
$$
\n
$$
P(A) = \frac{4}{8}
$$
\n
$$
P(B) =
$$



A: 1st person is a girl  
\nB: 2nd person is a girl  
\n
$$
P(A) = \frac{4}{8}
$$
\n
$$
P(B) = \frac{1}{2} \left( \begin{matrix} b_y & symmetry \end{matrix} \right)
$$
\n
$$
\left( \begin{matrix} or : sample space = 8.7 \end{matrix} \right)
$$
\n
$$
\left( \begin{matrix} or : each girl = 2nd, # outcomes = 7 \end{matrix} \right)
$$



A: 1st person is a girl B: 2nd person is a girl  $P(A) = \frac{4}{8}$  $P(B) = \frac{1}{2}$  (by symmetry) (or: sample space = 8.7 & for each girl=2nd, #outcomes=7)

 $P(B|A) =$ 



A: 1st person is a girl  
\n
$$
B: 2nd person is a girl\nP(A) =  $\frac{4}{8}$   
\nP(B) =  $\frac{1}{2}$  (by symmetry)  
\nfor: sample space = 8.7  
\n
$$
R = \frac{1}{8}
$$
 for each girl=2nd, # outcomes = 7  
\n
$$
P(B|A) = \frac{3}{7}
$$
$$



$$
P(car \otimes A) = P(A) = \frac{1}{3} = P(B) = P(C)
$$



$$
P(car \otimes A) = P(A) = \frac{1}{3} = P(B) = P(C)
$$
  
\nNow we know the car is not at B  
\nWhat is the probability if's at A?





$$
P(car \otimes A) = P(A) = \frac{1}{3} = P(B) = P(c)
$$





$$
P(car \otimes A) = P(A) = \frac{1}{3} = P(B) = P(c)
$$



$$
\frac{P(A|B)}{What we actually want is}
$$
\n
$$
P(A | (\underline{door B was opened} \cap \underline{we chose A}))
$$



BACK TO MONTY HALL : intuition



$$
P(car \otimes A) = P(A) = \frac{1}{3} = P(B) = P(c)
$$

BACK TO MONTY HALL : intuition  $P(car \otimes A) = P(A) = \frac{1}{2} = P(B) = P(C)$  $\boxed{\frac{1}{2}}$  $\begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$  $\begin{bmatrix} 0 & 0 \\ 0 & 1 \end{bmatrix}$  $w.l.o.g.$  guess A<br> $g \cdot g$  is shown.

$$
\begin{array}{ccccc}\n\frac{1}{3} & - & A & (car & \omega & A) \\
\hline\n\frac{1}{3} & - & B & & \\
\hline\n\frac{1}{3} & - & \frac{1}{3} & C & , then & B & recvaled & \rightarrow switch to win\n\end{array}
$$

BACK TO MONTY HALL : intuition  $P(car \otimes A) = P(A) = \frac{1}{2} = P(B) = P(C)$  $\boxed{\frac{1}{2}}$  $\begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$  $\begin{bmatrix} 0 & 0 \\ 0 & 1 \end{bmatrix}$  $w.l.o.g.$  guess A<br> $g \cdot g$  is shown.

$$
\begin{array}{cccc}\n\sqrt{3} & -4 & (car \otimes A) \\
& \sqrt{3} & -1 & 12 \\
\hline\n& \sqrt{3} & -1 & 12 \\
\hline\n& \sqrt{3} & -1 & 12 \\
\end{array}
$$



Back To Monry Hall: intuition  
\n
$$
\frac{A}{DU} = \frac{B}{\sqrt{14}} \frac{C}{DU}
$$
\n
$$
P(\text{car} \otimes A) = P(A) = \frac{1}{3} = P(B) = P(C)
$$
\n
$$
W.I.o.g. \text{ guess } A
$$
\n
$$
B \text{ is shown.}
$$
\n
$$
V.A \text{ B reveals}
$$
\n
$$
V.A \text{ B, then } C \text{ revealed } \rightarrow \text{suit.} + \text{ to win}
$$
\n
$$
V_{3} \rightarrow \text{ if } C \text{, then } B \text{ revealed } \rightarrow \text{suit.} + \text{ to win}
$$












$$
\frac{A}{\frac{dP}{dP}} = \frac{P[A | (\text{we chose } A \cap \text{door B was opened})]}{P(A | (\text{we chose } A \cap \text{door B was opened})}
$$

A	B	C
III	III	III

\n
$$
P(A | (\text{we chose } A \cap \text{door B was opened})]
$$
\n\n\n|\n Apply P(X|Y)=P(XAY) |\n II |\n
\n|\n IVY |\n IVY |\n IVY |\n
\n\n\n
$$
P(\text{we chose } A \cap \text{door B was opened})
$$
\n\n\n|\n Y |\n U |\n U |\n
\n|\n Y |\n U |\n U |\n
\n\n

A	B
III	CP
II	CP
II	CP

\n\n
$$
P(\omega e chose A \cap door B was opened)
$$
\n

\n\n $\omega \circ P = \frac{P[A \cap (we chose A) \cap door B was opened)}{P(we chose A) \cap door B was opened)}$ \n

\n\n $\omega \circ P = \frac{P[A \cap (we chose A) \cap door B was opened)}{P(we chose A) \cap door B was opened)}$ \n

\n\n $\omega \circ P = \frac{P[A \cap (we chose A) \cap door B was opened)}{P(we chose A) \cap door B was opened)}$ \n

So far,  
\n
$$
P[A | (\underbrace{\text{we chose } A \cap \text{door } B \text{ was opened}}_{A}]
$$
\n
$$
= \frac{P[ \text{door } B \text{ was opened} | (A \cap \text{we chose } A)] \cdot P(A \cap \text{we chose } A)}{P(\text{we chose } A \cap \text{door } B \text{ was opened})}
$$

So far,  
\n
$$
P[A \mid (\underbrace{\text{we chose A} \cap \text{door B was opened}}_{P(\text{we chose A} \cap \text{does A})}] \cdot P(A \cap \text{we chose A})
$$
\n
$$
= \frac{P[door B was opened | (A \cap \text{we chose A})] \cdot P(A \cap \text{we chose A})}{P(\text{we chose A} \cap \text{door B was opened})}
$$
\n
$$
= \frac{V_2 \cdot P(A) \cdot P(\text{we chose A})}{P(\text{we chose A} \cap \text{door B was opened})}
$$

So far,  
\n
$$
P[A \mid (\underbrace{\text{we chose } A \cap \text{door B was opened}}_{P(\text{we chose } A \cap \text{door B was opened}}) ]
$$
\n
$$
= \frac{P[\text{door B was opened} \mid (A \cap \text{we chose A})] \cdot P(A \cap \text{we chose A})}{P(\text{we chose } A \cap \text{door B was opened})}
$$
\n
$$
= \frac{\frac{1}{2} \cdot P(A) \cdot P(\text{we chose A})}{P(\text{we chose } A \cap \text{door B was opened})}
$$
\n
$$
= \frac{\frac{1}{2} \cdot (\frac{1}{3} \cdot \frac{1}{3})}{P(\text{we chose } A \cap \text{door B was opened})}
$$

So far,  
\n
$$
P[A | (\text{we chose A} \cap \text{door B was opened})
$$
\n
$$
= \frac{\frac{1}{2} \cdot \frac{1}{3} \cdot \frac{1}{3}}{P(\text{we chose A} \cap \text{door B was opened})}
$$

So far,  
\n
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{\text{20.14 not have}}]
$$
\n
$$
= \frac{V_1 \cdot V_3 \cdot V_3}{P(\text{we chose } A \cap \text{ door } B \text{ was opened})}
$$
\n
$$
= \frac{V_{18}}{\left[ P(\text{we chose } A \cap \text{ door } B \text{ was opened } \frac{\cap A}{\cap C} ) \right]}
$$
\n
$$
= \frac{V_{18}}{\left[ P(\text{we chose } A \cap \text{ door } B \text{ was opened } \frac{\cap A}{\cap C} ) \right]}
$$

The 2 terms in the denominator are "mutually exclusive", which means that they have no intersection. We can see that because one asks for A to happen, but the other asks for C to happen (among other things). In fact we should also include the term +P(we chose A AND door B was opened AND B) but this term is equal to zero.

Here's what is going on. There is an event, X. In our case,  $X =$  (we chose A AND door B was opened). We are interested in  $P(X)$ , as shown in orange above. We can say that  $P(X) = P(X \text{ and } A) + P(X \text{ and } B) + P(X \text{ and } C)$ , if A, B, C are mutually exclusive and cover all possibilities. Each of those 3 terms can have different values, but they will sum to P(X). Think of each term as a subset of X. (it makes sense that they are subsets of X, because they are more restrictive, and maintain all requirements of X). In fact not only are they subsets, but the 3 terms partition X. If X happens, then either (X and A) happens, or (X and B) happens, or (X and C) happens. In our case we have some additional information; X and B cannot happen.

So far,  
\n
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{\text{21} \cdot \text{1/3} \cdot \text{1/3}}
$$
\n
$$
= \frac{1/2 \cdot 1/3 \cdot 1/3}{P(\text{we chose } A \cap \text{ door } B \text{ was opened})}
$$
\n
$$
\left\{\n\begin{array}{r}\n\text{Cov}_1 | \text{not have} \\
\text{dov}_2 | \text{or B} \text{ or } B \text{ or } B \text{ is opened}\n\end{array}\n\right\}
$$
\n
$$
= \frac{1/18}{P(\text{we chose } A \cap \text{door } B \text{ was opened } \cap A)} \left[\n+ P(\text{we chose } A \cap \text{door } B \text{ was opened } \cap C)\n\right]
$$

 $P(\text{open} B(C \cap \text{choose} A)) = 1$ 

So far, 
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{1/2 \cdot 1/3 \cdot 1/3}
$$
  
=  $\frac{1/2 \cdot 1/3 \cdot 1/3}{P(\text{we chose } A \cap \text{ door } B \text{ was opened})}$   
 $\left[\frac{1/18}{\text{door } B \text{ was opened } \cap A}\right]$   
=  $\frac{1/18}{\left[P(\text{we chose } A \cap \text{door } B \text{ was opened } \cap A)\right]}$   
=  $\left[P(\text{we chose } A \cap \text{door } B \text{ was opened } \cap C)\right]$ 

 $P$  (open B  $(C \cap \text{choose A}) = 1$ <br>=  $\frac{P(C \cap \text{open B} \cap \text{choose A})}{P(C \cap \text{choose A})}$ 

So far,  
\n
$$
P[A \mid (\underbrace{\text{we chose A} \cap \text{ door B was opened}}_{\text{200C}})]
$$
\n
$$
= \frac{1/2 \cdot 1/3 \cdot 1/3}{P(\text{we chose A} \cap \text{door B was opened})}
$$
\n
$$
= \frac{1/18}{P(\text{we chose A} \cap \text{door B was opened} \cap A)}
$$
\n
$$
P(\text{open B} | (C \cap \text{choose A}) ) = \frac{1}{1} = \frac{1/18}{P(\text{we chose A} \cap \text{door B was opened} \cap C)}
$$
\n
$$
P(\text{open B} | (C \cap \text{choose A}) ) = \frac{1}{1} = \frac{1/18}{P(\text{we chose A} \cap \text{door B was opened} \cap A) + \frac{1}{3} \cdot \frac{1}{3}}
$$

$$
P[A | (we chose A \cap door B was opened)
$$
  
=  $\frac{1/18}{P(we chose A \cap door B was opened \cap A) + \frac{1}{3} \cdot \frac{1}{3}}$ 

So far,

$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{V18})
$$
\n
$$
= \frac{V_{18}}{P(\text{we chose } A \cap \text{ door } B \text{ was opened } \cap A) + \frac{1}{3} \cdot \frac{1}{3}}
$$
\n
$$
= \frac{V_{18}}{P[(A \cap \text{we chose } A) \cap \text{ door } B \text{ was opened } ] + V_{9}}
$$

 $\mathcal{L}^{\text{max}}(\mathbf{S})$  . The  $\mathcal{L}^{\text{max}}(\mathcal{L})$ 

So far,

So far,  
\n
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened})
$$
\n
$$
= \frac{1/18}{P(\text{we chose } A \cap \text{ door } B \text{ was opened } \cap A) + \frac{1}{3} \cdot \frac{1}{3}}
$$
\n
$$
= \frac{1/18}{P[(A \cap \text{we chose } A) \cap \text{ door } B \text{ was opened }] + \frac{1}{4}}
$$
\n
$$
= \frac{1/18}{P[\text{door } B \text{ was opened }] (A \cap \text{we chose } A)] \cdot P(A \cap \text{we chose } A) + \frac{1}{4}
$$

So far,  
\n
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{| / | B}
$$
\n
$$
= \frac{1/18}{P(\text{we chose } A \cap \text{ door } B \text{ was opened } \cap A) + \frac{1}{3} \cdot \frac{1}{3}}
$$
\n
$$
= \frac{1/18}{P[(A \cap \text{we chose } A) \cap \text{ door } B \text{ was opened}] + \frac{1}{4}}
$$
\n
$$
= \frac{1/18}{P[\text{door } B \text{ was opened } | (A \cap \text{we chose } A)] \cdot P(A \cap \text{we chose } A) + \frac{1}{4}
$$
\n
$$
= \frac{1/18}{P[\text{door } B \text{ was opened } | (A \cap \text{we chose } A)] \cdot P(A \cap \text{we chose } A) + \frac{1}{4}
$$

So far,  
\n
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{1/18})]
$$
\n
$$
= \frac{1/18}{P(\text{we chose } A \cap \text{ door } B \text{ was opened } \cap A) + \frac{1}{3} \cdot \frac{1}{3}}
$$
\n
$$
= \frac{1/18}{P[(A \cap \text{we chose } A) \cap \text{ door } B \text{ was opened}] + \frac{1}{4}}
$$
\n
$$
= \frac{1/18}{P[\text{door } B \text{ was opened} | (A \cap \text{we chose } A)] \cdot P(A \cap \text{we chose } A) + \frac{1}{4}
$$
\n
$$
= \frac{1/18}{P[\text{close} + \frac{1}{2} \cdot (\frac{1}{3} \cdot \frac{1}{3}) + \frac{1}{9}]}
$$

So far,  
\n
$$
P[A | (\underbrace{we \text{ chose } A \cap \text{ door } B \text{ was opened}}_{1/18}]
$$
\n
$$
= \frac{1/18}{P(\text{we chose } A \cap \text{ door } B \text{ was opened } \cap A) + \frac{1}{3} \cdot \frac{1}{3}}
$$
\n
$$
= \frac{1/18}{P[(A \cap \text{we chose } A) \cap \text{ door } B \text{ was opened}] + \frac{1}{4}}
$$
\n
$$
= \frac{1/18}{P[\text{door } B \text{ was opened} | (A \cap \text{we chose } A)] \cdot P(A \cap \text{we chose } A) + \frac{1}{4}
$$
\n
$$
= \frac{1/18}{P[\text{wole product} + \text{index} + \
$$

TESTING FOR A DISEASE · Suppose 1% of the population has a disease

· Suppose 1% of the population has a disease . there is a diagnostic test, that finds it 80% of the time<br>(assuming the subject has it)

- · Suppose 1% of the population has a disease
- . there is a diagnostic test, that finds it 80% of the time (assuming the subject has if) . the test also produces <u>false positives</u>, at a rate of 9.6% (you're fine, but the test says you're not)

· Suppose 1% of the population has a disease . there is a diagnostic test, that finds it 80% of the time<br>(assuming the subject has it)







 $P(disease|test:)= ?$ 











п





4 events					
Test ::	80%	90.6			
Test ::	20%	90.4%			
P(disease test ::)	P(disease \n $\frac{P(disease \n        \frac{P}{dset} ::)$	=			
P(1est ::)	=	P(disease \n $\frac{P}{dset} ::)$	=	P(1est ::)	P(1est ::)
P(1est ::)	=	P(1est ::)	=	P(1est ::)	
P(1est ::)	=	P(1est ::)	=	=	
P(1est ::)	=	=	=		
P(1est ::)	=	=	=		
P(1est ::)	=	=	=		
P(disease test ::)	=	=	=		

Bayes theorem
## $P(A \cap B) = P(A|B) \cdot P(B)$  $= P(B|A) \cdot P(A)$

 $P(A \cap B) = P(A|B) \cdot P(B)$  $= P(B|A) \cdot P(A)$  $\left(\begin{matrix} P(A|B) = P(B|A) \cdot \frac{P(A)}{P(B)} \end{matrix}\right)$ 



$$
P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\overline{A}) \cdot P(\overline{A})}
$$

$$
P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\overline{A}) \cdot P(\overline{A})}
$$

A: have disease  $B:$  test  $\therefore$ 

$$
P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\overline{A}) \cdot P(\overline{A})}
$$



$$
P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\overline{A}) \cdot P(\overline{A})}
$$
  
A: have disease |  

$$
1\%A \mid \text{qq}\% \overline{A} \qquad P(A) = 0.01
$$

$$
\mathsf{B}: \mathsf{test} \: : \:
$$

 $B/A = 80\%$   $B/\bar{A} = 9.6\%$ 

 $T(A) = 0.99$  $P(B|A) = 0.8$  $P(B|\overline{A}) = 0.096$ 

$$
P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\overline{A}) \cdot P(\overline{A})}
$$
  
A: have disease |  
1%A | 99% $\overline{A}$   $P(A) = 0.01$ 

 $B:$  test  $\therefore$ 

 $B/A = 80\%$   $B/\bar{A} = 9.6\%$ 

 $P(\bar{A}) = 0.99$  $P(B|A) = 0.8$  $P(B|\overline{A}) = 0.096$ 

 $0.8 \cdot 0.01$  $P(A|B) =$  $0.8 - 0.01 + 0.096 \cdot 0.99$